



### CRoNoS

### Workshop on Multivariate Data Analysis and Software

# A new method for curves clustering in general dependence models

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Sottile and Adelfio (UNIPA)

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### Introduction

- Clustering of effects curves in quantile regression models
  - Quantile regression (QR)
  - Quantile regression coefficients modeling (QRCM)
  - Clustering of effects curves method (CEC)
  - Computation
  - Simulations
  - Application



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  - Application
- 3 Conclusions

# Literature review

The problem of curves clustering is very complex and has been recently addressed in several fields:

- structural averaging in the context of computing an average (Kneip and Gasser, 1992);
- curves registrastion in statistics (Silverman, 1995; Ramsay and Li, 1998);
- time warping in engineering (Wang and Gasser, 1997)

# Literature review

### In statistics:

- Silverman, (1995) proposed a general approach, in which a target curve must satisfy a predefined criterion;
- Ramsay and Li, (1998) used a Procrustes fitting procedure (Gower, 1975) to provide maximal alignment to the target function;
- James, (2007) introduced a method for finding similarities between functions by equating the moments among all curves;
- Garcia-Escudero and Gordaliza, (2005) proposed a new approach based on the trimmed k-means Robust Curve Clustering;
- Adelfio et al., (2012) introduced a procedure to identify clusters of multivariate waveforms;
- Adelfio et al., (2016) focused on finding clusters of multidimensional curves with spatio-temporal structure.

# WHAT ABOUT CURVES CLUSTERING IN GENERAL DEPENDENCE MODELS?

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QR Koenker and Bassett Jr, (1978) and Koenker, (2005)

Let be *y* a response variable, and **x** a *q*-dimensional vector of covariates. We assume that  $Q(p | \mathbf{x}) = \mathbf{x}^T \beta(p)$  is the *p*-th quantile of *y*, given **x**. The vector of quantile regression coefficients,  $\beta(p)$ , can be estimated by

$$\hat{\beta}(\boldsymbol{p}) = \arg\min_{\boldsymbol{\beta}\in\mathcal{R}^q} \sum_{i=1}^n \omega_{\boldsymbol{p},i}(\boldsymbol{y}_i - \boldsymbol{x}_i^T \boldsymbol{\beta})$$

where  $\omega_{p,i} = \mathcal{I}(\mathbf{y}_i \leq \mathbf{x}_i^T \boldsymbol{\beta})$  and  $\mathcal{I}(\cdot)$  is the indicator function.

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### Issues

- quantiles are estimated one at the time
- the estimated coefficients are generally non-smooth functions of p

Quantile regression (QR)

### QR Koenker and Bassett Jr, (1978) and Koenker, (2005)



Estimated quantile regression coefficient and 95% pointwise confidence intervals (shaded area). The dotted line suggests a possible linear trend.

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### QRCM Frumento and Bottai, (2016)

A parametric approach to model the quantile function estimating the coefficients as functions of the order of the quantile  $p \in (0, 1)$ 

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$$Q(p \mid \boldsymbol{x}, \boldsymbol{\theta}) = \boldsymbol{x}^{T} \beta(p \mid \boldsymbol{\theta}),$$

- **x** is the model matrix  $(N \times q)$
- $\beta(p \mid \theta) = \theta b(p)$
- $\boldsymbol{b}(\boldsymbol{p}) = [1, b_1(\boldsymbol{p}), \dots, b_k(\boldsymbol{p})]^T$  is a set of (k + 1) known functions
- $\theta$  is the unknown parameter matrix

### QRCM Frumento and Bottai, (2016)



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# CEC method Sottile and Adelfio

### Aims

Our goal is to use the QRCM framework to answer two different questions:

 Univariate case. Given one response variable we estimate β<sub>1</sub>(p | θ),..., β<sub>q</sub>(p | θ) ⇒ the aim is to assess if these q curves, can be clustered based on similarities of effects

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- Univariate case. Given one response variable we estimate β<sub>1</sub>(p | θ),..., β<sub>q</sub>(p | θ) ⇒ the aim is to assess if these q curves, can be clustered based on similarities of effects
- Multivariate case. Given *m* response variables we estimate β<sub>11</sub>(*p* | θ),..., β<sub>mq</sub>(*p* | θ) ⇒ the **aim** is to assess if there are similar responses given covariates

### Our proposal

A new dissimilarity measure, that accounts both for the **shape** and for the **distance**:

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• the *shape* evaluated using its second derivative. Moreover, two different curves are similar in shape if, at any given point, the signs of the second derivatives are concordant;

$$d_{\text{shape}}^{\textit{ii'}}(\boldsymbol{p}) = \mathcal{I}(\text{sign}(\beta_{i}^{\prime\prime}(\boldsymbol{p} \mid \boldsymbol{\theta})) \times \text{sign}(\beta_{i^{\prime}}^{\prime\prime}(\boldsymbol{p} \mid \boldsymbol{\theta})) = \mathbf{1})$$

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 the distance between two curves evaluated as their differences with respect to other curves. Two curves are said close if their distance at any given point is lower than a fixed value;

$$d_{\text{distance}}^{ii'}(\boldsymbol{p}) = \mathcal{I}(|\beta_i(\boldsymbol{p} \mid \boldsymbol{\theta}) - \beta_{i'}(\boldsymbol{p} \mid \boldsymbol{\theta})| \le f(\alpha, \text{dist}(\boldsymbol{p})))$$

### The cut-off function

The  $f(\cdot, \cdot)$  function depends on a probability value  $\alpha$ , and on dist(p) that is, for each percentile, the distribution of all possible distances among curves. Therefore, the cut-off function selects the  $\alpha$ -th percentile of dist(p)

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### The role of $\alpha$

 $\alpha$  is the probability value, and it has a central role for finding hoogeneous clusters. Its choice depends on the goal of the analysis and has to be fixed by the researcher.

The new dissimilarity measure

$$d(i,i') = 1 - \int_0^1 \left[ d_{ ext{shape}}^{ii'}(oldsymbol{p}) \cdot d_{ ext{distance}}^{ij'}(oldsymbol{p}) 
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The product of the two measures is computed, to account for their concordance at each point.

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### Optimization

We implemented the above measure in the clustEff R package.

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# Computation

Pseudo code of the algorithm implemented in the  ${\tt clustEff}$  package

Step	Algorithm
1	fix the $\alpha$ -level and calculate dist( <b>p</b> ) for each $p \in (0, 1)$
2	the cut-off function $f(\cdot, \cdot)$ selects the percentiles of
	the distribution of dist( $m{p}$ ) used in $d_{ ext{distance}}^{ii'}(m{p})$
3	compute $d_{\text{shape}}^{ii'}(\boldsymbol{p}), d_{\text{distance}}^{ii'}(\boldsymbol{p})$ , and hence $d(i, i')$
4	apply a hierarchical clustering algorithm to the dissimilarity
	matrix in order to obtain the dendrogram
5	select the optimal number of clusters $l^*$ , unless it is known
	in advance
6	calculate goodness-of-fit measures

# Choice of the number of clusters

### Effects curves

$$\pi_{=}^{l} I \sum_{j=1}^{l} q_{j}^{-1} \sum_{i=1}^{q_{j}} \bigg\{ \int_{0}^{1} \mathcal{I}\bigg(\overline{\mathtt{LB}}_{j}(\boldsymbol{p}) \leq \beta_{j}^{i}(\boldsymbol{p} \mid \boldsymbol{\theta}) \leq \overline{\mathtt{UB}}_{j}(\boldsymbol{p})\bigg) d\boldsymbol{p} \bigg\},$$

The value  $l^*$  is identified by that partition for which  $\pi^{l} - \pi^{l+1}$  is minimized

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The value  $I^*$  is identified by that partition for which  $\pi^{I} - \pi^{I+1}$  is minimized

### General curves

$$\operatorname{dist}_{\operatorname{rel}}^{I} = I \sup_{j \in \{1, \dots, I\}} \bigg\{ q_{j}^{-1} \sum_{i=1}^{q_{j}} \int_{0}^{1} |\overline{\beta}_{j}(p) - \beta_{j}^{i}(p \mid \theta)| \ dp \bigg\}.$$

The value  $I^*$  is identified by that partition for which dist<sup>*l*</sup><sub>rel</sub> – dist<sup>*l*+1</sup><sub>rel</sub> is minimized

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# Simulation scenario 1

### **Clusters of effects**

We considered a multivariate scenario in which the general quantile function was defined by

$$Q(p \mid x, \theta) = \beta_0(p \mid \theta) + \beta_1(p \mid \theta)x$$

where  $x \sim \mathcal{U}(0, 5)$ .

# Simulation scenario 1

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where  $x \sim \mathcal{U}(0,5)$ .

We defined three quantile functions and generated 30 response variables, 10 for each of them, using polynomial trends, i.e.,

1.  $Q_1(p \mid x, \theta_1) = (1 + \phi(p)) + (.5 + .5p + p^2 + 2p^3)x$ 

2. 
$$Q_2(p \mid x, \theta_2) = (1 + \phi(p)) + (-3 + .5p + p^2 + .5p^3)x$$

3.  $Q_3(p \mid x, \theta_3) = (1 + \phi(p)) + (.3 - .5p - p^2 + 2p^3)x$ 

# Simulation scenario 1



Output of the proposed algorithm for one replicate. The left panel shows the dendrogram; the middle panel shows the 30 curves clustered in 3 groups; the right panel shows the boxplot of the average dissimilarity within each cluster.

# Simulation scenario 2

### Waveform clustering

We simulated 30 harmonic functions evaluated at a grid of size 1000

- 1.  $f(t) = \sin(3\pi t) (\times 10)$
- **2**.  $g(t) = \cos(3\pi t) (\times 13)$
- 3.  $h(t) = \sin(3\pi t) \cos(\pi t)$  (×5)
- **4**. l(t) = 0 (×2)

# Simulation scenario 2

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**4**. 
$$l(t) = 0 (\times 2)$$

A random error  $\epsilon_t \sim \mathcal{N}(0, \sigma_t)$  was added to each curve to define a segmented relation with multiple change-points, such as

$$\sigma_t = 4 \max(t - 0.2, 0) - 8 \max(t - 0.5, 0) + 4 \max(t - 0.8, 0)$$

# Simulation scenario 2



### Simulation scenario 2



Output of the proposed algorithm for one replicate. Upper panels show the 4 clusters; bottom-left panel shows the dendrogram; bottom-right panel shows the boxplot of the average dissimilarity within each cluster.

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### Methods

- funFem: a functional mixture model (Bouveyron and Brunet-Saumard, 2014)
- FPCA: a *k*-means algorithm based on the principal component rotation of data (Adelfio et al., 2011)

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### Mesasures

• Area
$$(I^*) = I^{*-1} \sum_{j=1}^{I^*} \left\{ q_j^{-1} \sum_{i=1}^{q_j} \int_0^1 \left( |\overline{\beta}_j(p) - \beta_j^i(p|\theta)| \right) dp$$
  
•  $\rho_{\text{dist}}(I^*) = I^{*-1} \sum_{j=1}^{I^*} \left\{ 1 - \left[ 2 \left( q_j(q_j-1) \right)^{-1} \sum_{i=1}^{q_j-1} \sum_{z>i}^{q_j} \rho_{iz} \right]^2 \right\}$ 

the average number of clusters *l*\*

# Simulations

Average area, average distance based on correlation ( $\rho_{dist}$ ) and average of the optimal number of discovered clusters ( $I^*$ ) using the three different algorithms (clustEff, funFEM and FPCA) and as benchmark measure the true partition of curves in the 100 runs. SD in brackets.

		True	clustEff	funFEM	FPCA
Sim 1	<i>I</i> *	3.00(0.00)	3.51(1.63)	5.06(0.98)	3.35(0.63)
	Area	0.216(0.091)	0.205(0.091)	0.178(0.079)	0.206(0.090)
	$ ho_{ m dist}$	0.010(0.015)	0.010(0.015)	0.008(0.008)	0.010(0.015)
Sim 2	<i>I</i> *	4.00(0.00)	4.23(0.95)	3.44(1.05)	3.83(0.38)
	Area	0.133(0.102)	0.130(0.099)	0.177(0.146)	0.142(0.116)
	$ ho_{\rm dist}$	0.441(0.177)	0.426(0.175)	0.436(0.203)	0.437(0.171)

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A study carried out in 1988-1991 in Northern Italy

- *N* = 2,045 subjects (51% Male and 49% Female)
- q = 9 (age, height, body mass index (BMI), sex, current smoking status, occupational exposure, cough, wheezing and asthma)

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### The model basis

- Intercept:  $b(p) = [1, \log(p), \log(1-p)]^T$
- Covariates: a shifted Legendre polynomials up to a 5<sup>th</sup> degree (Abramowitz and Stegun, 1964)

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Table 1: Average area, average correlation ( $\rho_{dist}$ ) and average of the optimal number of discovered clusters ( $I^*$ ) are compared across the three algorithm (clustEff, funFEM and FPCA).

	clustEff	funFEM	FPCA
<b>/</b> *	5	3	3
Area	0.004	0.039	0.039
$\rho_{\rm dist}$	0.197	0.710	0.710



The five clusters obtained applying the clustEff algorithm on the estimated quantile regression coefficients of inspiratory capacity dataset. Black solid lines are the mean curves; the dashed lines are the effects curves; the shaded areas are identified by the mean lower and upper bands within each cluster. The dotted line indicates the zero.

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- Results of two different simulation scenarios showed good performance of our proposal with respect of two competitors funFEM and FPCA;
- Results on the Inspiratory Capacity data showed a variable selection perspective.

### Thanks for the attention!!!

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